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THE EVALUATION OF THE COMPLETENESS OF DEATH
REGISTRATION IN THE PRESENCE OF HIGH NET OUT-MIGRATION.
THE CASE EXAMPLE OF MAURITIUS

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FOREWORD

As the developing countries continue in their efforts to improve their civil registration and vital statistics systems, it is important for them to be able to evaluate their progress. For estimating the completeness of death reporting, a number of indirect estimation techniques have been developed. These techniques assume a closed population, and so their usefulness for countries which have high rates of net in or out migration is a matter of concern. The author of this paper applies the techniques to the case of Mauritius, an island nation with high net out migration, to study the impact of this demographic condition on the estimates of completeness of death reporting.

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THE EVALUATION OF THE COMPLETENESS OF DEATH REGISTRATION IN THE PRESENCE
OF HIGH NET OUT-MIGRATION: THE CASE EXAMPLE OF MAURITIUS

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Introduction

Mauritius is a very small island of approximately 720 square miles. It is situated in the Indian Ocean, 500 miles east of Madagascar and over 1,000 miles from the southeast coast of Africa. In 1987, the population was about one million and is made up of three broad population groups: the Indo-Mauritians, the Sino-Mauritians and the 'general population'. The Indo-Mauritians comprising of immigrants from India and their descendants make up about two-thirds of the population, while the Sino-Mauritians comprising of immigrants from China and their descendants make up about 3 percent. The rest of the population is known as the 'general population' and is made up of persons from European or African or mixed decent.

In Mauritius, data on death and other vital statistics is collected by the Registrar General's Department, which by 1982, maintained 48 Civil Status Offices located throughout the country. Mauritius is known to have a tradition of very good birth and death registration statistics as well as censuses. According to the UN (1982,2):

"Research and evaluation by national and international organizations have shown that recent censuses in Mauritius have been quite accurate...Civil registration of vital events is compulsory and in the last few decades, with the exception of marriages, has been nearly complete."

The ascertainment of the completeness of death registration is important not only for its role in the estimation of intercensal populations but also in the construction of life tables which is used in many other demographic estimation procedures.

During the 1980s, several techniques were developed for the evaluation of census data and the evaluation of vital statistics data for the intercensal period. Some of these techniques were based on stable population theory or the more recent generalized stable population theory which relaxes the assumption of stability but retains the assumption of closure to migration. This paper is an exercise on the application of such techniques; namely, the estimation of the completeness of death registration. In Mauritius, vital registration is quite good but because of high population density, emigration is encouraged; as a result, net out-migration figures are high. The paper aims to investigate the behaviour of these techniques when applied to Mauritian data.

In indirect estimation, many of the techniques used are robust with respect to the violation of the assumptions of the technique. Since these relatively new techniques which make use of two censuses have potential to be used in many developing countries, it is worthwhile to investigate the degree of their robustness. In this case, Mauritian data is used as an example of data with high accuracy but one in which emigration rates are known to be high.

Literature review of relevant techniques

For the total population, the only source of error that one is concerned with is net under-enumeration error. Net misreporting error is no longer considered since all the misreported cases are included in the population total. Closely associated with this net under-enumeration error is the completeness of the death registration. The ascertainment of this completeness is important not only for its role in the estimation of intercensal populations and in the construction of life tables but also because recorded number of deaths are used as inputs in estimating the relative completeness of one census to another.

Various methods have been developed for estimating the completeness of death registration. Each of them assumes that the population is closed to migration and that ages at death were correctly reported. In addition to these, the following assume stability of the population:

- (1) Brass's sectional growth balance method (Brass, 1975)
- (2) Preston and Coale method (Preston, Coale et al, 1980)
- (3) Bourgeois-Pichat method (Preston, 1984)

The other methods which relax the assumption of stability are given below:

- (4) Forward projection method
- (5) The modified growth balance method (Martin, 1980)
- (6) Preston and Hill method (Preston and Hill, 1980)
- (7) Bennett and Horiuchi method (Bennett and Horiuchi, 1981)
- (8) United Nation method (United Nations, 1979).

Excellent summary of these techniques has been given by Preston (1984). The methods that do not assume stability are a generalisation of the methods assuming stability. One of these non-stable methods, the United Nations method, has not been much used and has shown to give erratic results (Preston, 1984).

In the modified growth balance method, Martin (1980) sought to adjust the Brass estimate of completeness, for mortality decline. The procedure requires the knowledge of the rate and duration of mortality decline. Estimation of these factors involves some circularity because the estimate of the mortality condition is the aim of the exercise in the first place.

The Preston and Hill method and the forward projection methods are related to the size of cohorts in two censuses. They become awkward to use for irregular intercensal periods as in Mauritius with an intercensal period of 11 years, 1972-1983. For such a purpose, one is left with the Bennett-Horiuchi method.

Method and material

The Bennett-Horiuchi method builds on the Preston and Coale (1980) method that assumes stability. Beside the advantage of relaxing the assumption of stability, the Bennett-Horiuchi technique can also be used for irregular intercensal periods. Instead, it makes the less restrictive assumption that the observed number of persons in each five-year interval is approximately equal to the corresponding number in a stable population inferred from the numbers at ages a and $a+5$ and the observed age-specific growth rates. This relationship is given by:

$$N(a) = \int_a^{\infty} D^*(x) \exp\left[\int_a^x r(u) du\right] dx \quad (1)$$

Where $N(a)$ is the number of persons aged a , $r(u)$ is the growth rate of the population aged u and $D^*(x)$ is the true number of deaths experienced by persons aged x . $D^*(x)$ is defined as:

$$D^*(x) = k(x)D(x) \quad (2)$$

Where $D(x)$ is the number of registered deaths to persons aged x and $k(x)$ is the inverse of the completeness of death registration, $c(x)$.

Using the concept of differential growth rates within a population, values of ${}_5N_a'$, the estimated number of persons in the age group a to $a+5$ are computed. Estimates of completeness are derived from the median of a series of age specific completeness of death (${}_5c_a$) which are ratios of estimated to observed populations (${}_5N_a' / {}_5N_a$). This median is then used to calculate an adjusted set of age-specific death rates and life expectancies for ages 5 and above (Bennett and Horiuchi, 1981).

Results

Using this Bennett and Horiuchi technique on Mauritian data, the results obtained are summarised in Tables 1 and 2 for males and females respectively.

In Table 1, eight of the age-specific estimates of completeness, ${}_5c_a$, are above unity. According to Preston (1984), one reason for the upward bias of ${}_5c_a$ is age over statement of deaths at higher ages.

Table 1
 ESTIMATED COMPLETENESS OF DEATH REGISTRATION AND ADJUSTED LIFE EXPECTANCY (APPLICATION OF BENNETT-HORIUCHI TECHNIQUE)
 Mauritius males, 1972-1983

| AGE | POPULATION | | GROWTH RATE | INTERCENSAL DEATHS | | COMPLETENESS (1) OF DEATH REGISTRATION | ADJUSTED LIFE TABLE (2) | |
|-------|------------|----------|----------------|--------------------|-------|--|-------------------------|-------------|
| | JUN 1972 | JUL 1983 | | NUMBER | RATE | | DEATH RATES | APPROX E(X) |
| 0-5 | 51312 | 56219 | 00824 | 7132 | 01198 | . | 01199 | ... |
| 5-10 | 60031 | 52533 | 01204 | 402 | 00065 | 835 | 00065 | 60.6 |
| 10-15 | 56205 | 48172 | 01392 | 389 | 00067 | 862 | 00067 | 55.8 |
| 15-20 | 50276 | 57480 | 01208 | 648 | 00109 | 909 | 00109 | 51.0 |
| 20-25 | 40190 | 53122 | 02517 | 742 | 00145 | 916 | 00145 | 46.3 |
| 25-30 | 26230 | 44746 | 04819 | 756 | 00199 | 991 | 00199 | 41.6 |
| 30-35 | 21189 | 39263 | 05565 | 816 | 00255 | 1 011 | 00255 | 37.0 |
| 35-40 | 20772 | 26978 | 02359 | 985 | 00375 | 947 | 00376 | 32.4 |
| 40-45 | 18328 | 19969 | 00774 | 1479 | 00698 | 981 | 00698 | 28.0 |
| 45-50 | 20159 | 19301 | 00392 | 2185 | 00999 | 996 | 01000 | 23.9 |
| 50-55 | 14549 | 16173 | 00955 | 3095 | 01820 | 1 034 | 01821 | 20.0 |
| 55-60 | 12629 | 17296 | 02837 | 3844 | 02347 | 1 057 | 02348 | 16.7 |
| 60-65 | 9222 | 11888 | 02291 | 4538 | 03910 | 1 003 | 03912 | 13.4 |
| 65-70 | 6073 | 8665 | 03207 | 4578 | 05694 | 1 042 | 05697 | 10.8 |
| 70-75 | 3677 | 5237 | 03191 | 4145 | 08522 | 1 048 | 08526 | 8.5 |
| 75-80 | 1686 | 2720 | 04315 | 2940 | 12387 | 1 072 | 12393 | 6.7 |
| 80-85 | 767 | 1159 | 03725 | 1712 | 16383 | 1 094 | 16391 | 5.3 |
| 85+ | 285 | 447 | 04061 | 1018 | 25734 | ... | 25746 | 3.9 |
| TOTAL | 413580 | 481368 | 01369 | 41404 | | | | |

(1) FOR CALCULATION PURPOSES, E(85) ASSUMED EQUAL TO 4.946

(2) BASED ON MEDIAN COMPLETENESS OF 1.000

Source: output from BENIR

Table 2
ESTIMATED COMPLETENESS OF DEATH REGISTRATION AND ADJUSTED LIFE EXPECTANCY (APPLICATION OF BENNETT-HORIUCHI TECHNIQUE)
Mauritius females, 1972-1983

| AGE | POPULATION | | GROWTH RATE | INTERCENSAL DEATHS | | COMPLETENESS (1) OF DEATH REGISTRATION | | ADJUSTED LIFE TABLE (2) | |
|-------|------------|----------|----------------|--------------------|-------|--|-------------|-------------------------|--|
| | JUN 1972 | JUL 1983 | | NUMBER | RATE | DEATH RATES | APPROX E(X) | | |
| 0- 5 | 50253 | 55214 | 00849 | 5861 | 01004 | | 01019 | ... | |
| 5-10 | 58681 | 51695 | - 01144 | 357 | 00058 | 810 | 00059 | 67 5 | |
| 10-15 | 55144 | 46701 | - 01499 | 311 | 00055 | 840 | 00056 | 62 7 | |
| 15-20 | 50484 | 56323 | 00987 | 586 | 00099 | 885 | 00101 | 57 9 | |
| 20-25 | 39270 | 52250 | 02577 | 668 | 00133 | 896 | 00135 | 53 1 | |
| 25-30 | 26999 | 44700 | 04549 | 604 | 00157 | 965 | 00159 | 48 5 | |
| 30-35 | 21518 | 38698 | 05295 | 588 | 00184 | 982 | 00187 | 43 9 | |
| 35-40 | 20492 | 27842 | 02766 | 611 | 00231 | 933 | 00234 | 39 2 | |
| 40-45 | 17700 | 20523 | 01335 | 716 | 00339 | 963 | 00344 | 34 7 | |
| 45-50 | 18461 | 19560 | 00522 | 905 | 00430 | 990 | 00436 | 30 2 | |
| 50-55 | 13223 | 16047 | 01746 | 1302 | 00806 | 1 041 | 00818 | 25 8 | |
| 55-60 | 12478 | 17526 | 03065 | 1822 | 01112 | 1 055 | 01128 | 21 8 | |
| 60-65 | 9584 | 12712 | 02548 | 2297 | 01878 | 995 | 01905 | 17 9 | |
| 65-70 | 6969 | 10006 | 03264 | 2676 | 02891 | 1 050 | 02934 | 14 5 | |
| 70-75 | 5047 | 7048 | 03013 | 3195 | 04833 | 1 065 | 04904 | 11 4 | |
| 75-80 | 3044 | 4581 | 03688 | 3123 | 07546 | 1 075 | 07656 | 8 8 | |
| 80-85 | 1727 | 2512 | 03381 | 2632 | 11401 | 1 077 | 11569 | 6 8 | |
| 85+ | 1095 | 1557 | 03176 | 2751 | 19009 | | 19288 | 5 2 | |
| TOTAL | 412169 | 485495 | 01477 | 31005 | | | | | |

(1) FOR CALCULATION PURPOSES, E(85) ASSUMED EQUAL TO 5 260

(2) BASED ON MEDIAN COMPLETENESS OF 986

Source: output from BEMER

As has been mentioned earlier, census age reporting was very good in Mauritius. It is not known whether this is the same for age reporting at death. Also, if the two population censuses are equally complete and if death registration is complete for all ages above 5, these series of 5c_x values will be more or less constant. In Table 1, the values of 5c_x range from 0.835 for age group 5-10 to 1.095 for age group 80-85. Indeed, from age group 65-70 onwards the values of 5c_x continuously rise.

The series of expectation of life values estimated by this technique also over estimate male mortality in Mauritius. From the Mauritius official life tables, the expectation of life at age 5, e^5_x is 61.17 and 61.49 in 1971 and 1982-84 respectively. It is expected that e_x values obtained from the Bennett-Horiuchi method should lie between those of the official life tables since the former refers to the intercensal period. Table 3 shows that this is the case for only three ages; 55, 60 and 65. For ages less than 55, the Bennett-Horiuchi method over estimates mortality while for ages greater than 65 it under estimates mortality.

Table 3

Comparison Between Expectation of Life Obtained From the Official Life tables and From the Bennet and Horiuchi Method. Mauritius Males, 1972-1983

| Age | Official e_x | Bennet-Horiuchi e_x | Official e_x |
|-----|----------------|-----------------------|----------------|
| | 1971-73 | 1972-1983 | 1982-1984 |
| 5 | 61.17 | 60.6 | 61.49 |
| 10 | 56.43 | 55.8 | 56.65 |
| 15 | 51.68 | 51.0 | 51.78 |
| 20 | 46.97 | 46.3 | 47.02 |
| 25 | 42.25 | 41.6 | 42.31 |
| 30 | 37.59 | 37.0 | 37.65 |
| 35 | 32.99 | 32.4 | 33.07 |
| 40 | 28.52 | 28.0 | 28.66 |
| 45 | 24.25 | 23.9 | 24.46 |
| 50 | 20.24 | 20.0 | 20.56 |
| 55 | 16.58 | 16.7 | 16.95 |
| 60 | 13.28 | 13.4 | 13.65 |
| 65 | 10.66 | 10.8 | 10.81 |
| 70 | 8.26 | 8.5 | 8.23 |
| 75 | 6.34 | 6.7 | 6.18 |
| 80 | 4.64 | 5.3 | 4.20 |

Source: Central Statistics Office (CSO), Table 1; Tables 5.13 and 5.14.

Quite unlike the estimated e_x values in Table 3 for males, those for females in Table 4 show remarkable consistency. Up to age 50, the estimated intercensal value obtained from the Bennett and Horiuchi method lie between those from the official life tables for the two successive censuses. However, after age 50, mortality seems to have been under estimated leading to higher values of life expectancy.

Table 4

Comparison Between Expectation of Life Obtained from the Official Life tables and from the Bennett and Horiuchi Method, Mauritius Females, 1972-1983

| Age | Official e_x | Bennet-Horiuchi e_x | Official e_x |
|-----|----------------|-----------------------|----------------|
| | 1971-73 | 1972-1983 | 1982-1984 |
| 5 | 65.93 | 67.5 | 68.18 |
| 10 | 61.27 | 62.7 | 63.32 |
| 15 | 56.45 | 57.1 | 58.46 |
| 20 | 51.76 | 53.1 | 53.72 |
| 25 | 47.21 | 48.5 | 48.99 |
| 30 | 42.66 | 43.9 | 44.24 |
| 35 | 38.24 | 39.2 | 39.51 |
| 40 | 33.75 | 34.7 | 34.84 |
| 45 | 29.38 | 30.2 | 30.31 |
| 50 | 25.06 | 25.8 | 25.88 |
| 55 | 21.00 | 21.8 | 21.61 |
| 60 | 17.18 | 17.9 | 17.66 |
| 65 | 13.77 | 14.5 | 14.03 |
| 70 | 10.51 | 11.4 | 10.68 |
| 75 | 7.95 | 8.8 | 7.89 |
| 80 | 5.62 | 6.8 | 5.23 |

Source: CSO, Table 2, Table 5.13 and 5.14.

Contrary to expectation, it is at those ages where mortality has been underestimated the most (above 65) that the completeness for under registration is consistently above unity.

Discussion

The over-estimation of mortality for ages under 55 could partly be due to genuine under reporting or due to the violation of the assumption of closure. One can be more certain of the later possibility as the population of Mauritius is not quite a closed one as can be seen from Table 5.

Table 5

Net Migration in Mauritius, 1973-1983

| Year | Arrivals A | Departure D | Net migration A - D |
|------|---------------|----------------|------------------------|
| 1973 | 101184 | 104697 | -3513 |
| 1974 | 109044 | 113362 | -4318 |
| 1975 | 117548 | 120703 | -3155 |
| 1976 | 139303 | 141148 | -1845 |
| 1977 | 153208 | 155653 | -2445 |
| 1978 | 161688 | 163733 | -2045 |
| 1979 | 182771 | 186864 | -4093 |
| 1980 | 163230 | 167269 | -4039 |
| 1981 | 168973 | 174376 | -5403 |
| 1982 | 166669 | 171991 | -5322 |
| 1983 | 177665 | 182005 | -4340 |

Source: CSO (1986) Extracted from Table 1.16

In the presence of net outward migration, emigrants appear as dead people under the assumption of a closed population. As a result, the estimate of completeness may be biased downward. This is summarised below:

From equation 2,

$$c(x) = \frac{l}{k(x)} = \frac{D(x)}{D^*(x)} \quad (3)$$

This expression shows that $c(x)$ would be low if $D(x)$ is low either due to omission in reporting of dead or due to age misstatement. Low values of $c(x)$ are, however, more plausible than values of $c(x)$ exceeding unity as there is very low chance that reported deaths could exceed true deaths. The reason for the upward bias of $c(x)$ could be because of large net out-migration and/or the use of higher age specific growth rate due to relative under-enumeration in the first census. These happen to be two of the limitations of this method which are also shared with other methods for estimating under-registration (Bennett and Horiuchi, 1981).

Summary and conclusion

The results of the application of the Bennett-Horiuchi technique showed a high degree of completeness of death registration in Mauritius. However, some results of the age specific completeness yielded improbable figures higher than unity. The problem could largely be due to out-migration. When out-migration is high and the population is assumed to be a closed one, the expected number of deaths may form a high proportion of reported deaths or may even exceed them. This may lead to an upward bias in the estimated values. However, the fact that the minimum and median values for completeness at specific ages ($c(x)$) are 0.835 and 1.000 respectively for males and 0.810 and 0.986 respectively for females confirms a very good registration system. According to Preston (1984), as a rough rule of thumb, a registration system that records 60 percent or more of deaths represents a very useful source of mortality information.

The presence of high out-migration in Mauritius makes the population an open one rather than being closed. This makes the application of methods for estimating completeness of death registration

slightly unsuited. One way to make the techniques more applicable is to adjust the data before applying them. It is suggested that if the extent of migration is significantly large compared to the number of deaths, prior adjustment of the data is necessary. Also, if the first census is relatively under-enumerated, the set of observed age specific growth rates may have to be deflated. This could be done by experimenting with different rates of growth to produce a sequence of completeness estimates that varies least with age. The former method is difficult in the absence of age specific migration rates while the latter may not be satisfactory (Preston, 1984). However, in many cases, data on age specific migration rates needed to do this are unavailable and the methods used for modifying the rates of growth are arbitrary. This makes adjustment of the data problematic.

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